Trevor Immelman of South Africa was the last golfer to make a hole in one at The Masters, acing the 16th hole in 2005. Eighteen holes in one have been scored during Masters play.



By Kevin Greer and Ron Coddington, USA TODAY Source: The Masters

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Suppose there are two populations: population 1, in which a proportion p_1 have a certain characteristic, and population 2, in which a proportion p_2 have a certain (possibly different) characteristic. We will use a sample of size n_1 from population 1, and n_2 from population 2 to estimate the difference $p_1 - p_2$.

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Specifically, if y_1 is the number having the population 1 characteristic in the n_1 items in sample 1, and if if y_2 is the number having the population 2 characteristic in the n_2 items in sample 2, then the sample proportion having the population 1 characteristic is $\hat{p}_1 = y_1/n_1$, and the sample proportion having the population 2 characteristic is $\hat{p}_2 = y_2/n_2$.

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A point estimator of $p_1 - p_2$ is $\hat{p}_1 - \hat{p}_2$.

The standard error of $\hat{p}_1 - \hat{p}_2$ is

$$\sqrt{rac{p_1(1-p_1)}{n_1}+rac{p_2(1-p_2)}{n_2}}$$

Further, for large n_1 and n_2 , the Central Limit Theorem ensures that $\hat{p}_1 - \hat{p}_2$ has approximately a normal distribution, so

$$rac{\hat{p}_1-\hat{p}_2-(p_1-p_2)}{\sqrt{rac{p_1(1-p_1)}{n_1}+rac{p_2(1-p_2)}{n_2}}}$$

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has approximately a N(0, 1) distribution.

Based on this, and on the fact that if n_1 and n_2 are large, then \hat{p}_1 and \hat{p}_2 are close to p_1 and p_2 , respectively, an approximate level L confidence interval for $p_1 - p_2$ has endpoints

$$\hat{p}_1 - \hat{p}_2 \pm z_{(1+L)/2} \sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}}$$

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However, by "fudging" the sample proportions in much the same way as we did in the one sample case, we can get an approximate interval that works well for all sample sizes.

Comparing Two Population Proportions Specifically, to compute the level L approximate score (or Agresti-Coull) interval, first compute the adjusted estimates of n_1 and n_2 :

$$\tilde{n}_1 = n_1 + 0.5 z_{(1+L)/2}^2, \ \tilde{n}_2 = n_2 + 0.5 z_{(1+L)/2}^2,$$

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and then the adjusted estimates of p_1 and p_2 :

$$ilde{p}_1 = rac{y_1 + 0.25 z_{(1+L)/2}^2}{ ilde{n}_1}, \ ilde{p}_2 = rac{y_2 + 0.25 z_{(1+L)/2}^2}{ ilde{n}_2}$$

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The approximate score interval for $p_1 - p_2$ is then given by the formula:

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In a recent survey on academic dishonesty 24 of the 200 female college students surveyed and 26 of the 100 male college students surveyed agreed or strongly agreed with the statement "Under some circumstances academic dishonesty is justified." With 95% confidence estimate the difference in the proportions p_f of all female and p_m of all male college students who agree or strongly agree with this statement.

1. The Scientific Goal:



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- The Statistical Model: Two independent binomials b(200, p_f), b(100, p_m).
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- 3. The Model Parameter(s) to Be Estimated: $p_f p_m$

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4. Point and Interval Estimates:

a. Point estimate:

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$$\tilde{n}_1 = 200 + 0.5 \cdot 1.96^2 = 201.9208, \ \tilde{n}_2 = 100 + 0.5 \cdot 1.96^2 = 101.9208$$

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The adjusted estimates of p_f and p_m are then

$$\tilde{p}_f = \frac{24 + 0.25 \cdot 1.96^2}{\tilde{n}_1} = 0.1236,$$

and

$$\tilde{p}_m = rac{26 + 0.25 \cdot 1.96^2}{\tilde{n}_2} = 0.2645.$$

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The approximate score interval for $p_f - p_m$ is then

 $0.1236 - 0.2645 \pm$

$$1.96\sqrt{\frac{0.1236(1-0.1236)}{201.9208} + \frac{0.2645(1-0.2645)}{101.9208}}$$
$$= (-0.2378, -0.0440)$$

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(SAS code here)

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5. Results and Interpretation:

 Results and Interpretation: With 95% confidence we estimate that the percentage of male college students who agree or strongly agree with the statement is between 4.4 and 23.78 percent greater than the corresponding percentage of female college students.

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• Population Versus Sample

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- Population Versus Sample
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- Population Versus Sample
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• Specific Estimation Problems:

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- Specific Estimation Problems:
 - o 1-Sample Mean, Known Variance

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$LDL \ Data$

Subject	Baseline	Follow-up	LDL Decrease
1	160.5	168.1	-7.6
2	195.3	181.4	13.9
3	181.7	154.6	27.1
4	175.1	160.3	14.8
5	198.3	192.0	6.3
6	215.5	173.5	42.0
7	227.9	186.2	41.7
8	201.7	183.2	18.5
9	161.5	130.3	31.2
10	189.0	165.0	24.0

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