

Appendix 6-1: Hypothesis Test Guide

This appendix provides a guide to the hypothesis tests covered in the chapter.

Hypothesis Tests for Means in the C+E Model: One Population

Case 1: Known Variances

Assumptions

1. The data are Y_1, Y_2, \dots, Y_n where $Y_j = \mu + \epsilon_j$.
2. Either
 - (a) n is large, or
 - (b) The ϵ_j are from a $N(0, \sigma^2)$ population.
3. σ^2 is known.

Formulas

Standardized Test Statistic¹: $Z = \frac{\bar{Y} - \mu_0}{\sigma(\bar{Y})}$.

Hypotheses:	$H_0 : \mu = \mu_0$ $H_{a-} : \mu < \mu_0$	$H_0 : \mu = \mu_0$ $H_{a+} : \mu > \mu_0$	$H_0 : \mu = \mu_0$ $H_{a\pm} : \mu \neq \mu_0$
p -value ² :	$p_- = P(N(0, 1) \leq z^*)$	$p^+ = P(N(0, 1) \geq z^*)$	$p_{\pm} = 2 \min(p_-, p^+)$

$${}^1\sigma(\bar{Y}) = \sqrt{\frac{\sigma^2}{n}}$$

² z^* is the observed value of Z , $P(N(0, 1) \leq z^*)$ is the proportion of a $N(0, 1)$ population less than z^* , and $P(N(0, 1) \geq z^*)$ is the proportion of a $N(0, 1)$ population greater than z^* .

Case 2: Unknown Variance, n Large

Assumptions 1 and 2a from Case 1 are assumed to hold. This case is treated exactly as Case 1 except that $\sigma(\bar{Y})$ is replaced by $\hat{\sigma}(\bar{Y}) = \sqrt{\frac{S^2}{n}}$ in the computation of the standardized test statistic Z , where S^2 is the sample variance computed from the data.

Case 3: Unknown Variance, n Small

Assumptions

1. The data are Y_1, Y_2, \dots, Y_n where $Y_j = \mu + \epsilon_j$.
2. The ϵ_j are from a $N(0, \sigma^2)$ population.
3. σ^2 is unknown.

Formulas

Standardized Test Statistic¹: $t = \frac{\bar{Y} - \mu_0}{\hat{\sigma}(\bar{Y})}$.

Hypotheses:	$H_0 : \mu = \mu_0$ $H_{a-} : \mu < \mu_0$	$H_0 : \mu = \mu_0$ $H_{a+} : \mu > \mu_0$	$H_0 : \mu = \mu_0$ $H_{a\pm} : \mu \neq \mu_0$
p -value ² :	$p_- = P(t_{n-1} \leq t^*)$	$p^+ = P(t_{n-1} \geq t^*)$	$p_{\pm} = 2 \min(p_-, p^+)$

$${}^1\hat{\sigma}(\bar{Y}) = \sqrt{\frac{S^2}{n}}$$

² t^* is the observed value of t , $P(t_{n-1} \leq t^*)$ is the proportion of a t_{n-1} population less than t^* , and $P(t_{n-1} \geq t^*)$ is the proportion of a t_{n-1} population greater than t^* .

Hypothesis Tests for the Proportion in the Binomial Model: One Population

Case 1: An Exact Test

Assumption

The datum is Y from a $b(n, p)$ population.

FormulasTest Statistic: Y .

Hypotheses:	$H_0 : p = p_0$ $H_{a-} : p < p_0$	$H_0 : p = p_0$ $H_{a+} : p > p_0$	$H_0 : p = p_0$ $H_{a\pm} : p \neq p_0$
p -value ¹ :	$p_- = P(b(n, p_0) \leq y^*)$	$p^+ = P(b(n, p_0) \geq y^*)$	p_{\pm} is given by Equation (6.4)

¹ y^* is the observed value of Y , $P(b(n, p_0) \leq y^*)$ is the proportion of a $b(n, p_0)$ population less than or equal to y^* , and $P(b(n, p_0) \geq y^*)$ is the proportion of a $b(n, p_0)$ population greater than or equal to y^* .

Case 2: An Approximate Test for n Large**Assumptions**

1. The datum is Y from a $b(n, p)$ population.
2. n is large: $Y \geq 10$ and $n - Y \geq 10$ is a good rule of thumb overall. If $0.3 \leq \hat{p} = Y/n \leq 0.7$, then $Y \geq 5$ and $n - Y \geq 5$ is a good rule of thumb.

FormulasStandardized Test Statistic: $Z = \frac{Y - np_0}{\sigma(Y)}$ where $\sigma(Y) = \sqrt{np_0(1 - p_0)}$.

Hypotheses:	$H_0 : p = p_0$ $H_{a-} : p < p_0$	$H_0 : p = p_0$ $H_{a+} : p > p_0$	$H_0 : p = p_0$ $H_{a\pm} : p \neq p_0$
p -value ¹ :	$p_- = P(N(0, 1) \leq z_l^*)$	$p^+ = P(N(0, 1) \geq z_u^*)$	$p_{\pm} = 2 \min(p_-, p^+)$

¹ $z_l^* = \frac{y^* - np_0(1 - p_0) + 0.5}{\sigma(Y)}$, $z_u^* = \frac{y^* - np_0(1 - p_0) - 0.5}{\sigma(Y)}$, where y^* is the observed value of Y . $P(N(0, 1) \leq z_l^*)$ is the proportion of a $N(0, 1)$ population less than z_l^* , and $P(N(0, 1) \geq z_u^*)$ is the proportion of a $N(0, 1)$ population greater than z_u^* .

Hypothesis Tests for Differences in Means in the C+E Model: Two Independent Populations**Case 1: Known Variances****Assumptions**

1. The data are

$$Y_{1,1}, Y_{1,2}, \dots, Y_{1,n_1}, \text{ where } Y_{1,j} = \mu_1 + \epsilon_{1,j}, \text{ (population 1)}$$

$$Y_{2,1}, Y_{2,2}, \dots, Y_{2,n_2}, \text{ where } Y_{2,j} = \mu_2 + \epsilon_{2,j}, \text{ (population 2).}$$
2. The two populations are independent.
3. Either
 - (a) n_1 and n_2 are large, or
 - (b) The $\epsilon_{1,j}$ are from a $N(0, \sigma_1^2)$ population, and the $\epsilon_{2,j}$ are from a $N(0, \sigma_2^2)$ population.
4. σ_1^2 and σ_2^2 are known.

FormulasStandardized Test Statistic¹: $Z = \frac{\bar{Y}_1 - \bar{Y}_2 - \delta_0}{\sigma(\bar{Y}_1 - \bar{Y}_2)}$.

Hypotheses:	$H_0 : \mu_1 - \mu_2 = \delta_0$ $H_{a-} : \mu_1 - \mu_2 < \delta_0$	$H_0 : \mu_1 - \mu_2 = \delta_0$ $H_{a+} : \mu_1 - \mu_2 > \delta_0$	$H_0 : \mu_1 - \mu_2 = \delta_0$ $H_{a\pm} : \mu_1 - \mu_2 \neq \delta_0$
p -value ² :	$p_- = P(N(0, 1) \leq z^*)$	$p^+ = P(N(0, 1) \geq z^*)$	$p_{\pm} = 2 \min(p_-, p^+)$

$${}^1\sigma(\bar{Y}_1 - \bar{Y}_2) = \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}$$

² z^* is the observed value of Z , $P(N(0, 1) \leq z^*)$ is the proportion of a $N(0, 1)$ population less than z^* , and $P(N(0, 1) \geq z^*)$ is the proportion of a $N(0, 1)$ population greater than z^* .

Case 2: Unknown Variances, n_1 and n_2 Large

Assumptions 1, 2 and 3a are assumed to hold. This case is treated exactly as Case 1 except that $\sigma(\bar{Y}_1 - \bar{Y}_2)$ is replaced by $\hat{\sigma}(\bar{Y}_1 - \bar{Y}_2) = \sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}$ in the computation of the standardized test statistic Z , where S_1^2 and S_2^2 are the sample variances computed from the data from populations 1 and 2 respectively.

Case 3: Variances Unknown, but Assumed to Be Equal, n_1 and n_2 Not Both Large

Assumptions

1. The data are

$$Y_{1,1}, Y_{1,2}, \dots, Y_{1,n_1}, \text{ where } Y_{1,j} = \mu_1 + \epsilon_{1,j}, \text{ (population 1)}$$

$$Y_{2,1}, Y_{2,2}, \dots, Y_{2,n_2}, \text{ where } Y_{2,j} = \mu_2 + \epsilon_{2,j}, \text{ (population 2).}$$

2. The two populations are independent.
3. The $\epsilon_{1,j}$ are from a $N(0, \sigma_1^2)$ population, and the $\epsilon_{2,j}$ are from a $N(0, \sigma_2^2)$ population.
4. σ_1^2 and σ_2^2 are unknown, but are assumed to be equal.

Formulas

Standardized Test Statistic¹: $t^{(p)} = \frac{\bar{Y}_1 - \bar{Y}_2 - \delta_0}{\hat{\sigma}_p(\bar{Y}_1 - \bar{Y}_2)}$.

Hypotheses:	$H_0 : \mu_1 - \mu_2 = \delta_0$ $H_{a-} : \mu_1 - \mu_2 < \delta_0$	$H_0 : \mu_1 - \mu_2 = \delta_0$ $H_{a+} : \mu_1 - \mu_2 > \delta_0$	$H_0 : \mu_1 - \mu_2 = \delta_0$ $H_{a\pm} : \mu_1 - \mu_2 \neq \delta_0$
p -value ² :	$p_- = P(t_{n_1+n_2-2} \leq t^{(p)*})$	$p^+ = P(t_{n_1+n_2-2} \geq t^{(p)*})$	$p_{\pm} = 2 \min(p_-, p^+)$

¹ $\hat{\sigma}_p(\bar{Y}_1 - \bar{Y}_2) = \sqrt{S_p^2 \left(\frac{1}{n_1} + \frac{1}{n_2}\right)}$, where $S_p^2 = \frac{(n_1-1)S_1^2 + (n_2-1)S_2^2}{n_1+n_2-2}$, and S_1^2 and S_2^2 are the sample variances computed from the data from populations 1 and 2 respectively.

² $t^{(p)*}$ is the observed value of $t^{(p)}$, $P(t_{n_1+n_2-2} \leq t^{(p)*})$ is the proportion of a $t_{n_1+n_2-2}$ population less than $t^{(p)*}$, and $P(t_{n_1+n_2-2} \geq t^{(p)*})$ is the proportion of a $t_{n_1+n_2-2}$ population greater than $t^{(p)*}$.

Case 4: Variances Unknown, and Not Assumed to Be Equal, n_1 and n_2 Not Both Large

Assumptions

1. The data are

$$Y_{1,1}, Y_{1,2}, \dots, Y_{1,n_1}, \text{ where } Y_{1,j} = \mu_1 + \epsilon_{1,j}, \text{ (population 1)}$$

$$Y_{2,1}, Y_{2,2}, \dots, Y_{2,n_2}, \text{ where } Y_{2,j} = \mu_2 + \epsilon_{2,j}, \text{ (population 2).}$$

2. The two populations are independent.
3. The $\epsilon_{1,j}$ are from a $N(0, \sigma_1^2)$ population, and the $\epsilon_{2,j}$ are from a $N(0, \sigma_2^2)$ population.
4. σ_1^2 and σ_2^2 are unknown, and are not assumed to be equal.

Formulas

Standardized Test Statistic¹: $t^{(ap)} = \frac{\bar{Y}_1 - \bar{Y}_2 - \delta_0}{\hat{\sigma}(\bar{Y}_1 - \bar{Y}_2)}$.

Hypotheses:	$H_0 : \mu_1 - \mu_2 = \delta_0$ $H_{a-} : \mu_1 - \mu_2 < \delta_0$	$H_0 : \mu_1 - \mu_2 = \delta_0$ $H_{a+} : \mu_1 - \mu_2 > \delta_0$	$H_0 : \mu_1 - \mu_2 = \delta_0$ $H_{a\pm} : \mu_1 - \mu_2 \neq \delta_0$
p -value ² :	$p_- = P(t_{\nu} \leq t^{(ap)*})$	$p^+ = P(t_{\nu} \geq t^{(ap)*})$	$p_{\pm} = 2 \min(p_-, p^+)$

¹ $\hat{\sigma}(\bar{Y}_1 - \bar{Y}_2) = \sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}$, where S_1 and S_2 are the sample standard deviations computed from the data from populations 1 and 2 respectively. Under H_0 , the distribution of $t^{(ap)}$ is approximately t_{ν} where the degrees of

freedom ν is taken as the largest integer less than or equal to $\left[\left(\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2} \right)^2 \right] / \left[\frac{\left(\frac{S_1^2}{n_1} \right)^2}{n_1-1} + \frac{\left(\frac{S_2^2}{n_2} \right)^2}{n_2-1} \right]$

² $t^{(ap)*}$ is the observed value of $t^{(ap)}$, $P(t_{\nu} \leq t^{(ap)*})$ is the proportion of a t_{ν} population less than $t^{(ap)*}$, and $P(t_{\nu} \geq t^{(ap)*})$ is the proportion of a t_{ν} population greater than $t^{(ap)*}$.

Hypothesis Tests for Proportions in the Binomial Model: Two Independent Populations

Case 1: A Test of the Equality of Two Proportions

Assumptions

1. The data are Y_1 from a $b(n_1, p_1)$ population and Y_2 from a $b(n_2, p_2)$ population.
2. The two populations are independent.
3. n_1 and n_2 are large: $Y_i \geq 10$ and $n_i - Y_i \geq 10$, $i = 1, 2$, is a good rule of thumb overall. If $0.3 \leq \hat{p}_i = Y_i/n_i \leq 0.7$, then $Y_i \geq 5$ and $n_i - Y_i \geq 5$, $i = 1, 2$, is a good rule of thumb.

Formulas

Standardized Test Statistic¹: $Z = \frac{\hat{p}_1 - \hat{p}_2}{\hat{\sigma}_0(\hat{p}_1 - \hat{p}_2)}$.

Hypotheses:	$H_0 : p_1 - p_2 = 0$ $H_{a-} : p_1 - p_2 < 0$	$H_0 : p_1 - p_2 = 0$ $H_{a+} : p_1 - p_2 > 0$	$H_0 : p_1 - p_2 = 0$ $H_{a\pm} : p_1 - p_2 \neq 0$
p -value ² :	$p_- = P(N(0, 1) \leq z^*)$	$p^+ = P(N(0, 1) \geq z^*)$	$p_{\pm} = 2 \min(p_-, p^+)$

$${}^1\hat{\sigma}_0(\hat{p}_1 - \hat{p}_2) = \sqrt{\hat{p}(1 - \hat{p}) \left(\frac{1}{n_1} + \frac{1}{n_2} \right)}, \text{ where } \hat{p} = \frac{Y_1 + Y_2}{n_1 + n_2}.$$

² z^* is the observed value of Z , $P(N(0, 1) \leq z^*)$ is the proportion of a $N(0, 1)$ population less than z^* , and $P(N(0, 1) \geq z^*)$ is the proportion of a $N(0, 1)$ population greater than z^* .

Case 2: A Test of the General Difference of Two Proportions

Assumptions

1. The data are Y_1 from a $b(n_1, p_1)$ population and Y_2 from a $b(n_2, p_2)$ population.
2. The two populations are independent.
3. n_1 and n_2 are large: a guideline is given in Case 1.

Formulas

Test Statistic¹: $Z = \frac{\hat{p}_1 - \hat{p}_2 - \delta_0}{\hat{\sigma}(\hat{p}_1 - \hat{p}_2)}$.

Hypotheses ² :	$H_0 : p_1 - p_2 = \delta_0$ $H_{a-} : p_1 - p_2 < \delta_0$	$H_0 : p_1 - p_2 = \delta_0$ $H_{a+} : p_1 - p_2 > \delta_0$	$H_0 : p_1 - p_2 = \delta_0$ $H_{a\pm} : p_1 - p_2 \neq \delta_0$
p -value ³ :	$p_- = P(N(0, 1) \leq z^*)$	$p^+ = P(N(0, 1) \geq z^*)$	$p_{\pm} = 2 \min(p_-, p^+)$

$${}^1\hat{\sigma}(\hat{p}_1 - \hat{p}_2) = \sqrt{\frac{\hat{p}_1(1 - \hat{p}_1)}{n_1} + \frac{\hat{p}_2(1 - \hat{p}_2)}{n_2}}, \text{ where } \hat{p}_1 = Y_1/n_1 \text{ and } \hat{p}_2 = Y_2/n_2.$$

² $\delta_0 \neq 0$.

³ z^* is the observed value of Z , $P(N(0, 1) \leq z^*)$ is the proportion of a $N(0, 1)$ population less than z^* , and $P(N(0, 1) \geq z^*)$ is the proportion of a $N(0, 1)$ population greater than z^* .