

Why Covariance and Correlation Measure Linearity

Lecture XIX Supplement

Covariance and Correlation

Covariance

$$\text{Cov}(X, Y) \stackrel{\text{Definition}}{=} \mathbb{E}[(X - \mu_X)(Y - \mu_Y)] \stackrel{\text{Shortcut}}{=} \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y].$$

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Correlation

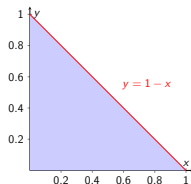
The correlation between X and Y is

$$\rho_{X,Y} = \frac{\text{Cov}(X, Y)}{\sigma_X \sigma_Y},$$

where σ_X, σ_Y are the standard deviations.

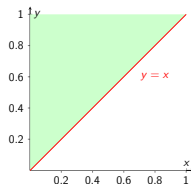
Note that $|\rho_{X,Y}| \leq 1$.

Two Examples from Last Class



(a) $f(x, y) = 2, (x, y) \in T$:
Negative-covariance.

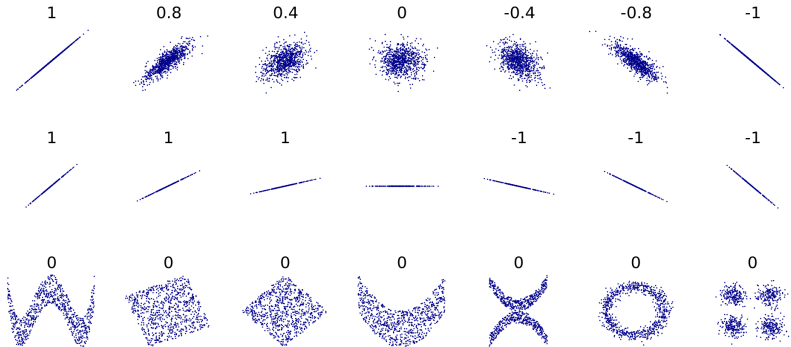
$$\begin{aligned}\text{Cov}(X, Y) &= -\frac{1}{36} \\ \sigma_X &= \sigma_Y = \frac{1}{\sqrt{18}} \\ \rho_{X,Y} &= \frac{\text{Cov}(X, Y)}{\sigma_X \sigma_Y} = -\frac{1}{2}\end{aligned}$$



(b) $f(x, y) = 2, (x, y) \in T$:
Positive-covariance.

$$\begin{aligned}\text{Cov}(X, Y) &= \frac{1}{36} \\ \sigma_X &= \sigma_Y = \frac{1}{\sqrt{18}} \\ \rho_{X,Y} &= \frac{\text{Cov}(X, Y)}{\sigma_X \sigma_Y} = \frac{1}{2}\end{aligned}$$

Some Example Distributions and Their Correlation Values



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This least-squares perspective shows that covariance and correlation quantify *linear dependence* — no more, no less.

Motivation

We know covariance appears in formulas, but *why* does it measure **linear** dependence (and it also fails to measure nonlinearity)?

Idea: Ask for the best linear predictor of Y in terms of X .

We try to fit a line:

$$\hat{Y} = aX + b$$

that minimizes the mean squared error, defined by

$$L(a, b) := \mathbb{E}[(Y - aX - b)^2].$$

Step 1: Differentiate with respect to b

$$L(a, b) = \mathbb{E}[(Y - aX - b)^2].$$

Differentiate:

$$\frac{\partial}{\partial b} L(a, b) = \mathbb{E}[2(Y - aX - b)(-1)].$$

So

$$\frac{\partial}{\partial b} L(a, b) = -2(\mathbb{E}[Y] - a\mathbb{E}[X] - b).$$

Set to zero:

$$b = \mu_Y - a\mu_X.$$

Step 2: Differentiate with respect to a

$$L(a, b) = \mathbb{E}[(Y - aX - b)^2].$$

Differentiate:

$$\frac{\partial}{\partial a} L(a, b) = \mathbb{E}[2(Y - aX - b)(-X)].$$

So

$$\frac{\partial}{\partial a} L(a, b) = -2 \mathbb{E}[X(Y - aX - b)].$$

Set to zero:

$$\mathbb{E}[XY] - a\mathbb{E}[X^2] - b\mu_X = 0.$$

Step 3: Substitute $b = \mu_Y - a\mu_X$

$$\mathbb{E}[XY] - a\mathbb{E}[X^2] - (\mu_Y - a\mu_X)\mu_X = 0.$$

Simplify:

$$\mathbb{E}[XY] - \mu_X\mu_Y - a(\mathbb{E}[X^2] - \mu_X^2) = 0.$$

Recognize covariance and variance:

$$\text{Cov}(X, Y) - a \text{Var}(X) = 0.$$

Step 4: The optimal slope and intercept

So the optimal coefficients are

$$a^* = \frac{\text{Cov}(X, Y)}{\text{Var}(X)}, \quad b^* = \mu_Y - a^* \mu_X.$$

Interpretation:

- ▶ If $\text{Cov}(X, Y) = 0$, then $a^* = 0$: the best linear predictor is $\hat{Y} = \mu_Y$.
- ▶ Otherwise, the sign of covariance determines the direction of the slope.

Step 5: Connection to correlation

Rewrite slope as

$$a^* = \rho_{X,Y} \frac{\sigma_Y}{\sigma_X}.$$

- ▶ $\rho_{X,Y}$ (correlation) tells us the *strength and direction* of the linear relation.
- ▶ The ratio σ_Y/σ_X rescales units.

By Cauchy–Schwarz inequality (you will learn this at one point in your life):

$$|\rho_{X,Y}| \leq 1,$$

with equality iff Y is an exact linear function of X .

Takeaway

Key Point

Covariance and correlation arise naturally from the *population least-squares problem*.

- ▶ Covariance = raw numerator in the optimal slope.
- ▶ Correlation = standardized measure of linearity, bounded between -1 and 1 .
- ▶ If correlation is 0 , best linear predictor is a flat line $\hat{Y} = \mu_Y$.

Therefore: Covariance and correlation quantify *linear dependence*, not arbitrary dependence.