



---

Ferguson Distributions Via Polya Urn Schemes

Author(s): David Blackwell and James B. MacQueen

Reviewed work(s):

Source: *The Annals of Statistics*, Vol. 1, No. 2 (Mar., 1973), pp. 353-355

Published by: [Institute of Mathematical Statistics](#)

Stable URL: <http://www.jstor.org/stable/2958020>

Accessed: 08/04/2012 13:18

---

Your use of the JSTOR archive indicates your acceptance of the Terms & Conditions of Use, available at <http://www.jstor.org/page/info/about/policies/terms.jsp>

JSTOR is a not-for-profit service that helps scholars, researchers, and students discover, use, and build upon a wide range of content in a trusted digital archive. We use information technology and tools to increase productivity and facilitate new forms of scholarship. For more information about JSTOR, please contact support@jstor.org.



*Institute of Mathematical Statistics* is collaborating with JSTOR to digitize, preserve and extend access to *The Annals of Statistics*.

<http://www.jstor.org>

## FERGUSON DISTRIBUTIONS VIA PÓLYA URN SCHEMES

BY DAVID BLACKWELL<sup>1</sup> AND JAMES B. MACQUEEN<sup>2</sup>

*University of California, Berkeley*

The Polya urn scheme is extended by allowing a continuum of colors. For the extended scheme, the distribution of colors after  $n$  draws is shown to converge as  $n \rightarrow \infty$  to a limiting discrete distribution  $\mu^*$ . The distribution of  $\mu^*$  is shown to be one introduced by Ferguson and, given  $\mu^*$ , the colors drawn from the urn are shown to be independent with distribution  $\mu^*$ .

Let  $\mu$  be any finite positive measure on (the Borel sets of) a complete separable metric space  $X$ . We shall say that a random probability measure  $\mu^*$  on  $X$  has a *Ferguson distribution with parameter  $\mu$*  if for every finite partition  $(B_1, \dots, B_r)$  of  $X$  the vector  $\mu^*(B_1), \dots, \mu^*(B_r)$  has a Dirichlet distribution with parameter  $\mu(B_1), \dots, \mu(B_r)$  (when  $\mu(B_i) = 0$ , this means  $\mu^*(B_i) = 0$  with probability 1). Ferguson [3] has shown that, for any  $\mu$ , Ferguson  $\mu^*$  exist and when used as prior distributions yield Bayesian counterparts to well-known classical nonparametric tests. He also shows that  $\mu^*$  is a.s. discrete. His approach involves a rather deep study of the gamma process.

One of us [1] has given a different and perhaps simpler proof that Ferguson priors concentrate on discrete distributions. In this note we give still a third approach to Ferguson distributions, exploiting their connection with generalized Pólya urn schemes.

We shall say that a sequence  $\{X_n, n \geq 1\}$  of random variables with values in  $X$  is a *Pólya sequence with parameter  $\mu$*  if for every  $B \subset X$

$$(1) \quad P(X_1 \in B) = \mu(B)/\mu(X)$$

and

$$(2) \quad P\{X_{n+1} \in B \mid X_1, \dots, X_n\} = \mu_n(B)/\mu_n(X),$$

where  $\mu_n = \mu + \sum_{i=1}^n \delta(X_i)$  and  $\delta(x)$  denotes the unit measure concentrating at  $x$ . Note that, for finite  $X$ , the sequence  $\{X_n\}$  represents the results of successive draws from an urn where initially the urn has  $\mu(x)$  balls of color  $x$  and, after each draw, the ball drawn is replaced and another ball of its same color is added to the urn. Note also that, without the restriction to finite  $X$ , for any (Borel measurable) function  $\phi$  on  $X$ , the sequence  $\{\phi(X_n)\}$  is a Pólya sequence with parameter  $\phi\mu$ , where  $\phi\mu(A) = \mu\{\phi \in A\}$ .

We now describe the connections between Pólya sequences and Ferguson distributions.

---

Received August 13, 1971.

<sup>1</sup> Research sponsored by the Air Force Office of Scientific Research, Office of Aerospace Research, USAF, under Grant AFOSR-71-2100.

<sup>2</sup> Visiting professor of Psychology and Statistics.

**THEOREM.** *Let  $\{X_n\}$  be a Pólya sequence with parameter  $\mu$ . Then*

- (a)  $m_n = \mu_n/\mu_n(X)$  converges with probability 1 as  $n \rightarrow \infty$  to a limiting discrete measure  $\mu^*$ ,
- (b)  $\mu^*$  has a Ferguson distribution with parameter  $\mu$  and
- (c) given  $\mu^*$ , the variables  $X_1, X_2, \dots$  are independent with distribution  $\mu^*$ .

**PROOF.** Suppose first that  $X$  is finite, say  $X = \{1, 2, \dots, r\}$ . Let  $\mu^*, \{X_n\}$  be variables whose joint distribution is defined by (b) and (c). If  $\pi_n$  is the empirical distribution of  $X_1, \dots, X_n$ , it follows from the strong law of large numbers that  $\pi_n \rightarrow \mu^*$  with probability 1 as  $n \rightarrow \infty$ . Since

$$m_n = (\mu + n\pi_n)/(\mu(X) + n),$$

(a) follows. It remains to show that  $\{X_n\}$  is a Pólya sequence with parameter  $\mu$ , which is equivalent to

$$(3) \quad P(A) = \prod_x \mu(x)^{n(x)} / \mu(X)^{[n]}$$

where  $A = \{X_1 = x_1, \dots, X_n = x_n\}$ ,  $n(x)$  denotes the number of  $i$  with  $x_i = x$  and  $a^{[k]} = a(a + 1)\dots(a + k - 1)$ .

Since  $P(A | \mu^*) = \prod_x \mu^*(x)^{n(x)}$ , we get

$$(4) \quad P(A) = E \prod_x \mu^*(x)^{n(x)}.$$

That the right sides of (3) and (4) are equal is a standard formula [2] for the moments of Dirichlet distributions.

For general  $X$ , let  $\{X_n\}$  be a Pólya sequence with parameter  $\mu$ , let  $I_j$  be the indicator of the event that  $X_j$  is different from all  $X_i$  with  $i < j$  and define

$$\begin{aligned} f_{nj} &= I_j m_n(X_j) & \text{for } 1 \leq j \leq n, \\ f_{nj} &= 0 & \text{for } j > n. \end{aligned}$$

We show that

$$(5) \quad \text{with probability 1, } f_{nj} \text{ converges as } n \rightarrow \infty,$$

say to  $f_j^*$  and

$$(6) \quad \sum_j f_j^* = 1 \text{ with probability 1.}$$

Part (a) of the Theorem, with  $\mu^*$  defined by

$$\mu^*(B) = \sum_{x_j \in B} f_j^*$$

is an easy consequence of (5) and (6) since, for any  $B$ , we have, writing

$$(7) \quad \begin{aligned} s_{nr} &= \sum_{1 \leq j \leq r; x_j \in B} f_{nj}, & t_{nr} &= \sum_{1 \leq j \leq r} f_{nj} \\ s_{nr} &\leq m_r(B) \leq s_{nr} + (1 - t_{nr}) & & \text{for } 1 \leq r \leq n, \end{aligned}$$

so that, letting first  $n \rightarrow \infty$ , then  $r \rightarrow \infty$  we obtain (a).

To get (5) and (6), fix  $r$  and define

$$\begin{aligned} U_n &= j & \text{if } 1 \leq j \leq r \text{ and } I_j = 1 \text{ and } X_{r+n} = X_j \\ &= 0 & \text{otherwise.} \end{aligned}$$

Given  $X_1, \dots, X_r$ , the sequence  $\{U_n\}$  is a Pólya sequence on  $\{0, 1, \dots, r\}$  with parameter  $\mu'$  defined by

$$\begin{aligned} \mu'(j) &= \mu_r(X)f_{rj} && \text{for } 1 \leq j \leq r, \\ \mu'(0) &= \mu_r(X) - \sum_{j=1}^r \mu'(j), \end{aligned}$$

and the sequence  $m_n'$  associated with  $\{U_n\}$  satisfies

$$(8) \quad m_n'(j) = f_{r+n,j} \quad \text{for } 1 \leq j \leq r$$

and

$$(9) \quad m_n'(0) = 1 - \sum_{j=1}^r f_{r+n,j}.$$

We apply the finite case of our Theorem to  $\{U_n\}$ . From (8) and part (a) of the Theorem we get (5), and from (9) and part (b) of the Theorem we conclude

$$(10) \quad E(1 - \sum_{j=1}^r f_j^* | X_1, \dots, X_r) = \frac{\mu'(0)}{\mu_r(X)} \leq \frac{\mu(X)}{\mu(X) + r}.$$

Taking expectation in (10) and letting  $r \rightarrow \infty$  gives  $E(1 - \sum_{j=1}^{\infty} f_j^*) = 0$ , and (6) follows.

Parts (b) and (c) are now easy consequences of the finite case. For any finite partition  $B_1, \dots, B_r$  of  $X$ , define  $\phi$  on  $X$  by  $\phi = i$  on  $B_i$ , so that  $\{\phi(X_n)\}$  is a Pólya sequence with parameter  $\phi\mu$ . We conclude that the limit of  $(m_n(B_1), \dots, m_n(B_r))$ , already identified as  $(\mu^*(B_1), \dots, \mu^*(B_r))$ , has a Dirichlet distribution with parameter  $\mu(B_1), \dots, \mu(B_r)$ , establishing (b). For (c), let  $\{\phi_i\}$  be a sequence of functions on  $X$ , each with finitely many values, such that, if  $\mathcal{F}_i$  is the (finite) field of  $X$ -sets determined by  $\phi_i$ , we have  $\mathcal{F}_{i+1} \supset \mathcal{F}_i$  and the Borel field determined by  $\mathcal{F} = \bigcup \mathcal{F}_i$  consists of all Borel sets. Part (c) of the finite case of our Theorem, applied to  $\{\phi_j(X_n)\}$ , yields

(c') given  $\phi_j \mu^*$ , the sequence  $\{\phi_i(X_n)\}$  is independent with distribution  $\phi_i \mu^*$  for  $i \leq j$ .

Letting  $j \rightarrow \infty$ , we get

(c'') given  $\mu^*$ , the sequence  $\{\phi_i(X_n)\}$  is independent with distribution  $\phi_i \mu^*$  for all  $i$ .

Since  $\{\phi_i(X_n)\}$  is independent with distribution  $\phi_i \mu^*$  for all  $i$  implies  $\{X_n\}$  is independent with distribution  $\mu^*$ , part (c) follows from (c''), completing the proof.

REFERENCES

[1] BLACKWELL, DAVID (1973). Discreteness of Ferguson selections. *Ann. Statist.* **1** 356-358.  
 [2] DEGROOT, MORRIS H. (1970). *Optimal Statistical Decisions*. McGraw-Hill, New York.  
 [3] FERGUSON, THOMAS S. (1972). A Bayesian analysis of some nonparametric problems. *Ann. Statist.* **1** 209-230.